Validity of the Human System Audit Transformed Leadership Short Scale (HSA-TFL) in four european countries

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ABSTRACT
This study aimed to clarify the validity of the short scale of Transformational Leadership used by the Human System Audit (short HSA-TFL). The need of today’s enterprises for combined assessment of transformational leadership and quality-related performance in wider contexts requires short instruments based on scientific research. Convergent, construct and criterion validity of the short HSA-TFL were analyzed. Comparison of the short HSA-TFL with the Multifactor Leadership Questionnaire (MLQ-5X) showed high convergent validity. Exploratory factor analysis with hospital workers in Spain (N=625) showed the single factor structure of the Spanish version of the HSA-TFL. Confirmatory factor analysis using three further samples of hospital workers (N= 776) from different european countries confirmed a single factor. As regards criterion validity, the results indicated that the short HSA-TFL is positively related in all four countries to subjective performance. In sum, the results provide empirical evidence for the validity of the short HSA-TFL scale.

Key words authors
Construct validity, short scale, transformational leadership.

Key words plus
Human relations, leadership, psychometry.


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subjetivo (validez de criterio). En conclusión, la versión breve del HSA-TFL es válida para el análisis del liderazgo transformacional.

Palabras clave autores
Validación de constructo, escala corta, liderazgo transformacional.

Palabras clave descriptores
Relaciones humanas, liderazgo, psicometría.

The concept of transformational leadership is of particular relevance to the enterprises of today’s complex world (Jung, Chow & Wu, 2003). The most well-known operationalization of this concept, the Multifactor Leadership Questionnaire (MLQ), was developed by Bass (1985). However, despite the considerable interest in Bass’ transformational leadership (TFL) concept, some aspects have been subject to intense debate. One of the most important of these is the ambiguity concerning the differentiation and number of sub-dimensions of Bass’ transformational leadership model. Furthermore, very strong relationships have been reported between the transformational leadership factors (Avolio, Bass & Jung, 1999; Carless, 1998; Tejeda, Scandura & Pillai, 2001). Enterprises today have an increased need for combined assessment of leadership and quality-related aspects of performance (Lowe, Kroeck & Sivasubramaniam, 1996; Molero, Cuadrado, Navas & Morales, 2007). To meet these needs it is necessary to develop a valid instrument that is easy and quick to apply (Carless, Wearing & Mann, 2000), based on scientific research (Felfe, 2006) and which leads to recommendations for performance-related development and intervention for leadership in a wider organizational context. This was the main aim of the present study. Following these ideas we consider transformational leadership theory in the interest of today’s organizations and conclude that there is a need for a TFL short scale. The short HSA-TFL scale is part of the organizational behavior conceptual framework of the Human System Audit (HSA) (Quijano, 2006).

Transformational leadership theory in the interest of today’s organizations

Bass (1985) based his theory of transformational leadership on Burns’s (1978) conceptualization, with several modifications. Following Bass (1985), the four dimensions of transformational leadership are charisma or idealized influence, inspirational motivation, intellectual stimulation, and individualized consideration.

After more than 25 years of accumulated research evidence the effectiveness of transformational leadership is acknowledged throughout the literature. Empirical studies show that transformational facets have a stronger relationship to success and to both individual and organizational outcome criteria (Zhu, Chew & Spangler, 2005) than do transactional scales (Lowe et al., 1996). Several meta-analyses have also provided evidence for the criterion-related validity of transformational and charismatic leadership (Dumdum, Lowe & Avolio, 2002; Fuller, Patterson, Hester & Stringer, 1996; Judge & Piccolo, 2004), which consistently showed a positive impact on both subjective (Lowe et al., 1996) and objective (Barling, Weber & Kelloway, 1996; Geyer & Steyrer, 1998) performance criteria. Transformational leadership behavior has been empirically linked to increased employee (e.g., Podsakoff, MacKenzie, Moorman & Fetter, 1990) and job satisfaction (Nemanicha & Seller, 2007), organizational commitment (e.g. Bycio, Hackett & Allen, 1995), supervisor-rated performance, extra effort (e.g. Seltzer & Bass, 1990), overall employee (e.g. Yammarino, Spangler & Bass, 1993) and unit performance (Bass, Avolio, Jung & Berson, 2003) and organizational effectiveness (Lowe et al., 1996). Given the need of today’s organizations for combined assessment of quality-related aspects of performance (Quijano, Navarro, Yepes, Berger & Romeo, 2008) the importance for organizations of transformational leadership behavior (Molero et al., 2007) in wider practical and cultural contexts seems obvious. The well-known MLQ (Bass, 1985) is rather long and this makes it difficult to use in practical circumstances. Moreover, from the scientific point of view, some authors (Burchell & Marsh, 1992) also report that refusal to participate in assessments could be influenced by the length of the questionnaire, and it thus seems necessary to have a reduced set of items to measure transformational leadership. These practical reasons can be consi-
dered to outweigh the psychometric advantages of a longer scale (Muck, Hell & Gosling, 2007).

Mixed empirical support for the transformational leadership model as developed by Bass (1985, 1995)

Despite the effectiveness of transformational leadership its proposed structure has proved highly controversial (e.g., Deluga & Souza, 1991; Rafferty & Griffin, 2004; Tejeda et al., 2001; Tepper & Percy, 1994; Tracey & Hinkin, 1998). Past research on transformational leadership placed the emphasis on a more differentiated model of this style of leadership (Bass & Avolio, 1995), although high correlations between the transformational scales have often been addressed in the literature (Avolio et al., 1999; Bycio et al., 1995; Carless, 1998; Den Hartog, Van Muijen & Koopman, 1997).

As a result of the mixed empirical support for a differentiated transformational model, authors such as Carless (1998) and Tepper and Percy (1994) have argued that research should examine the higher-order factors of transformational leadership rather than the individual components of the model. Given this mixed empirical support for the four-factor-model of transformational leadership it is worth investigating its structure when using the concept for performance-related assessment.

Overview of the present study

Taking these findings into account the objective of the present study was to clarify empirically the evidence for the validity of the Spanish version of the short scale for transformational leadership (short HSA-TFL-ES) developed as part of the Human System Audit framework (Quijano et al., 2008). Evidence of validity was based on different sources. First, to analyze convergent validity, we compared the short HSA-TFL to the MLQ. Secondly, we analyzed the dimensionality of the scale by conducting an exploratory factor analysis using a Spanish sample. To confirm the structure of the scale, confirmatory factor analysis was applied using three samples (one from the United Kingdom, one from Poland and one from Portugal). Following Den Hartog, House, Hanges, Ruiz-Quintanilla and Dorfman (1999) we expect for the HSA-TFL short scale for transformation leadership a universal construct across cultures. The criterion validity of this short HSA-TFL was subsequently assessed using subjective criteria variables (satisfaction, commitment and identification). In line with the results of other studies (Bycio et al., 1995; Lowe et al., 1996; Nemanicha & Seller, 2007; Podsakoff et al., 1990) our hypothesis is that transformational leadership is positively related in all four European countries to satisfaction, commitment and identification. We also based evidence for validity on internal consistency.

Methods

Participants

Data were collected from four samples. The first sample, which was used for exploratory factor analysis, consisted of employees from a medium-sized hospital in Spain. From a total of 1094 target participants, responses were received from 625 employees (response rate = 57.16 %). These employees assessed the transformational leadership style of their direct leader or supervisor by responding to the short HSA-TFL questionnaire. The professional occupations of these participants were distributed as follows: 17 % were doctors, 28 % nurses, 3.5 % other qualified staff, 18.2 % nursing assistants and 19 % non-healthcare staff (19 %), with the remaining 12.2 % not falling under any of these categories. No response was received from 2.1 %. Of the total, 9.9 % identified themselves as managers.

For the confirmatory factor analysis we used three samples of 776 hospital workers, each one from a different European country (UK, Poland and Portugal) and comprising 50 doctors, 555 nurses and 171 other workers, all of whom responded to the short HSA-TFL questionnaire. As before, the employees assessed the transformational leadership style of their direct leader or supervisor.
Measures

Transformational leadership

The Human System Audit short scale for transformational leadership (HSA-TFL-ES)

The transformational leadership short scale (HSA-TFL-ES) was developed as part of the Human System Audit (Quijano et al., 2008) framework for quality related assessment of human resources and processes. It is designed to evaluate participants’ perceptions of their supervisors’ transformational leadership. The items reflect charisma, inspirational motivation, intellectual stimulation and individualized consideration. The short HSA-TFL-ES was developed from the long version of the HSA-TFL-ES using principles such as semantic heterogeneity of the items, content validity, and a high level of explained variance of the items as regards the conceptual dimensions. The content validity of the Spanish version of the short-scale was considered adequate. A panel of five expert judges with extensive active research and applied content experience judged the selected items at the beginning of the development process. The applied scale consists of eight items, with two for each sub-dimension. The items are rated on a 5-point Likert scale ranging from 0 ( = definitely do not agree) to 4 ( = frequently, most of the time, completely agree). The scale is also available in Polish, Portuguese and English versions, which were developed through a careful process of translation and back-translation into Portuguese, English and Polish by native speakers who were experts in the field of organizational psychology. The content experts in each culture did not see a reason to modify or add items to the existing Spanish version.

Multifactor Leadership Questionnaire
MLQ (5X) Spanish version

In order to test the convergent validity of the leadership instrument (the Spanish version of the HSA-TFL) we used the subordinates’ version of the Spanish Multifactor Leadership Questionnaire MLQ (5X), validated by Molero (1994). This comprises questions that are specifically phrased for subordinates to evaluate their leader, and items are scored on a Likert-type scale from 0 = definitely not to 4 = frequently, most of the time.

Subjective criteria

The study included three subjective measures that were used to analyze the criterion validity of the HSA-TFL-ES. Respondents were asked to rate their identification with the organization, their organizational commitment (Identification/commitment Inventory ICI) and their satisfaction with the leader and the organization. These scales ranged from 1 ( = definitely do not agree) to 5 ( = completely agree) in all four participating countries. The items for the subjective criteria were again translated and back-
translated by native speakers, who were experts in the field of organizational psychology.

**Identification/Commitment Inventory (HSA-ICI)**

The questionnaires for organizational commitment and identification were developed by Quijano, Navarro and Cornejo (2000). The scale for Organizational Commitment (OC) is based on the concepts of Meyer and Allen (1991) and Mael and Ashforth (1992). The inventory represents a holistic model of OC, integrating attitudinal and behavioral perspectives. Organizational Identification (OI) is conceptually based on the ideas of O'Reilly and Chatman (1986), Franklin (1975) and Mael and Ashforth (1992). OI is considered as a type of link between employees and organization, which implies cognition, affect and desire. The questionnaire consists of twenty items. Its development and good psychometric properties (Cronbach’s Alpha of 0.941) are described by Romeo, Yepes, Berger, Guàrdia, and Castro (2011).

**Satisfaction Inventory (HSA-SI)**

The satisfaction scale was developed by Yepes-Baldó, Romeo, Berger, Díaz de Quijano, and Gómez-Benito (under revision). It is based on the concepts of Meliá and Peiró (1989) and Meliá, Peiró, and Calatayud (1986). It comprises eight items that cover satisfaction with the support and recognition of superiors, with the relationship with colleagues, with learning and professional development, with the physical working conditions, with job security and stability as regards the future, and, finally, satisfaction with salary and social benefits. The scale presents with an alpha of Cronbach of 0.0771 a good internal consistency, similar to that obtained by the abovementioned and other authors (Hunt, Osborn & Schuler, 1978; Schriesheim, 1979; Schriesheim, Kinicki & Schriesheim, 1979), as well as good validity. One factor is explaining 38.86% of the variance. The scales’ concept and psychometric properties are reported by Yepes-Baldó et al. (under revision).

**Procedure**

The questionnaires were administered to employees of hospitals in Spain, Portugal, the United Kingdom and Poland over a two-week period. Full anonymity was assured and participants answered voluntarily. Once completed the questionnaires were returned directly to the researchers.

**Statistical Analysis**

To analyze the convergent validity of the short HSA-TFL-ES the MLQ was used to correlate the short HSA-TFL-ES with the Spanish version of the Multifactor Leadership Questionnaire (MLQ-5X). The convergent validity of the short HSA-TFL-ES was also assessed by testing both scales for any differences between the means of the composite scores of transformational leadership, using the MLQ and the short HSA-TFL-ES in a Spanish sample of 105 hospital workers. Results were compared using the Wilcoxon signed-rank test for paired samples, which was chosen because the distribution of scores for the HSA short-scale did not follow a normal curve and the data are ordinal.

To explore the structure of the HSA-TFL-ES an exploratory factor analysis with a Spanish sample of 625 hospital workers was conducted. To assess the adequacy of the sample the Kaiser-Meyer-Olkin index (KMO) and Bartlett’s test of sphericity (BTS) were calculated. The principal components extraction method was used to identify the factors that explained the most variance and correlations. The following criteria were established: KMO should be greater than 0.5; BTS was set at $p < 0.05$; the item loadings should be greater than 0.40; the explained variance of the first factor should be more than 40%; and the item-scale correlation should be more than 0.35.

To test the explored structure and the factor validity of the leadership instrument, a confirmatory factor analysis (CFA) was conducted with a further sample of 776 hospital workers from Portugal, the United Kingdom and Poland. This CFA was performed using the LISREL program, version 8.6 (Jöreskog & Sörbom, 2006a) in order to test whether the
The weighted least squares (WLS) method was used for estimation, as recommended by Jöreskog and Aish (1993) for ordinal data. In a previous step, covariance matrices and asymptotic covariance matrices were obtained by using Prelis 2, version 2.50 (Jöreskog & Sörbom, 2006b). Following the recommendations of Kaplan (2000) and MacCallum and Austin (2000), various alternative criteria were used to evaluate the model’s overall goodness of fit. The following indices were used: a) \( \chi^2 \) likelihood ratio statistic; b) root mean square error of approximation (RMSEA) and its corresponding 90% confidence intervals; c) goodness of fit index (GFI); d) non-normed fit index (NNFI); and e) comparative fit index (CFI). Indicators of a good fit are that \( \chi^2 \) is not significant; that the GFI, NNFI and CFI have values above 0.90; and that the RMSEA value does not exceed 0.08.

To analyze the criterion validity of the HSA-TFL-ES its relationship with the subjective performance criteria was studied using a correlation design. Subordinates from all four countries were asked to rate the following subjective performance indicators: organizational commitment, organizational identification and satisfaction. The hypothesis that the HSA-TFL is positively related to subjective performance in different countries was tested.

Internal consistency was evaluated as a measure of the reliability of the HSA-TFL-ES. This was done by calculating Cronbach’s alpha, which was considered to be the optimal method for determining the internal consistency as it takes into account the degree of covariance between the test items. The following criterion was established: the value of Cronbach’s alpha should be at least 0.8. This analysis was performed with 625 Spanish hospital workers.

**Results**

**Dimensionality**

The short HSA-TFL-ES was first studied through exploratory factor analysis, and we therefore used the items as indicators. When exploring the dimensions of the leadership short scale the following results were obtained for the Spanish sample: adequacy of the data for factor analysis was high, with a Kaiser-Meyer-Olkin index of 0.936 and Bartlett’s test of sphericity = 4005.127 (\( p < 0.001 \)). Loadings between 0.822 and 0.890 indicated that all the items were good indicators of the construct. Principal components analysis yielded one factor, which explained 72.78% of the variance. The scree plot confirmed a one-factor solution, as well as the high item inter-correlations and item-scale correlations (see Table 2).

The next step involved testing whether the unidimensionality found in the Spanish sample would be replicated in the other European samples, comparing a one-factor with a four-factor model. The factor loadings were all significant (\( p < 0.01 \)) and notably high, thus illustrating the relevance of the corresponding item in measuring the construct and indicating that the chosen indicators provide a reliable measure of it. With respect to fit indices the values were as follows: \( \chi^2 (20) = 267.63, p < 0.01 \); RMSEA = 0.08 (CI = 0.07-0.09); GFI = 0.99; NNFI = 0.97; CFI = 0.98. All the fit indices, apart from \( \chi^2 \), reached the values established as satisfactory, since the RMSEA is 0.08 and the GFI, NNFI and CFI are all above 0.95. Given the large sample size used in this study, the \( \chi^2 \) value may lead to the rejection of acceptable models due to excessive power (Hayduk, 1996); therefore, we rely only on the alternative fit indices.

We also tested whether the structure of a four-factor model as proposed by Bass could be confirmed. However, the four-factor model presents some Heywood cases, with two correlations greater than 1 between factors. This led us to propose a more complex model that maintains the four factors and also specifies a second-order factor that is subject to the four primary factors. However, we obtained the same results with this second model: the fit indices reached good values but the model still presented some Heywood cases, this time with two standardized gamma values greater than 1s. These findings led us to accept the unidimensional model as the best representation of the data.
**Table 2**
Factor loadings means, standard deviations, intercorrelations and item-scale correlation for the items of the HSA-TFL-ES at individual level

<table>
<thead>
<tr>
<th>Factor loading</th>
<th>Mean</th>
<th>SD</th>
<th>I1</th>
<th>I2</th>
<th>I3</th>
<th>I4</th>
<th>I5</th>
<th>I6</th>
<th>I7</th>
<th>I8</th>
<th>Item-scale correlation</th>
</tr>
</thead>
<tbody>
<tr>
<td>I1</td>
<td>0.832</td>
<td>3.28</td>
<td>1.047</td>
<td>1.000</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.779</td>
</tr>
<tr>
<td>I2</td>
<td>0.859</td>
<td>2.77</td>
<td>1.091</td>
<td>0.709**</td>
<td>1.000</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.812</td>
</tr>
<tr>
<td>I3</td>
<td>0.855</td>
<td>3.17</td>
<td>1.004</td>
<td>0.707**</td>
<td>0.708**</td>
<td>1.000</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.806</td>
</tr>
<tr>
<td>I4</td>
<td>0.822</td>
<td>2.94</td>
<td>1.114</td>
<td>0.611**</td>
<td>0.672**</td>
<td>0.625**</td>
<td>1.000</td>
<td></td>
<td></td>
<td></td>
<td>0.767</td>
</tr>
<tr>
<td>I5</td>
<td>0.846</td>
<td>3.06</td>
<td>1.062</td>
<td>0.652**</td>
<td>0.674**</td>
<td>0.656**</td>
<td>0.718**</td>
<td>1.000</td>
<td></td>
<td></td>
<td>0.797</td>
</tr>
<tr>
<td>I6</td>
<td>0.851</td>
<td>3.07</td>
<td>1.031</td>
<td>0.629**</td>
<td>0.642**</td>
<td>0.677**</td>
<td>0.647**</td>
<td>0.684**</td>
<td>1.000</td>
<td></td>
<td>0.801</td>
</tr>
<tr>
<td>I7</td>
<td>0.867</td>
<td>3.02</td>
<td>1.046</td>
<td>0.659**</td>
<td>0.665**</td>
<td>0.675**</td>
<td>0.650**</td>
<td>0.676**</td>
<td>0.808**</td>
<td>1.000</td>
<td>0.820</td>
</tr>
<tr>
<td>I8</td>
<td>0.890</td>
<td>2.76</td>
<td>1.053</td>
<td>0.669**</td>
<td>0.785**</td>
<td>0.737**</td>
<td>0.694**</td>
<td>0.705**</td>
<td>0.708**</td>
<td>0.758**</td>
<td>1.000</td>
</tr>
</tbody>
</table>

(n = 625). I1: “I have trust in his/her ability to overcome any obstacle”, I2: “He/She develops ways of motivating us”, I3: “I feel proud to work with him/her”, I4: “He/She is concerned with training those who need it”, I5: “He/She gives advice to those who need it”, I6: “He/She gets us to rely on reasoning and evidence to solve problems”, I7: “He/She promotes the use of intelligence to overcome obstacles”, I8: “He/She presents things through an approach that stimulates me

Source: Own work

**Relationships with other variables**

The high correlation between the HSA-TFL-ES and the MLQ (5X) ($r = 0.84$, $p < 0.001$) supports the convergent validity between these constructs. Additionally, the comparison of average transformational leadership scores as measured by the HSA-TFL-ES and the MLQ reveals that MLQ scores ($X = 1.91$) were almost equal to the average transformational leadership scores as measured by the HSA-TFL-ES ($X = 1.92$) (both scale measures between 0 and 4). The Wilcoxon signed-rank test revealed no significant differences between the average ratings of transformational leadership as measured by the MLQ and the HSA short-scale ($z = -2.84$, $p > 0.05$, n.s.). Table 3 shows that the transformational leadership style measured by the short HSA-TFL-ES was significantly associated with subjective performance (organizational commitment, organizational identification and satisfaction) in all countries.

**Table 3**
Correlations for leadership and subjective performance indicators

<table>
<thead>
<tr>
<th>Leadership</th>
<th>Spain</th>
<th>Portugal</th>
<th>Poland</th>
<th>U. Kingdom</th>
</tr>
</thead>
<tbody>
<tr>
<td>Identification</td>
<td>0.515**</td>
<td>0.328**</td>
<td>0.507**</td>
<td>0.351**</td>
</tr>
<tr>
<td>Commitment</td>
<td>0.599**</td>
<td>0.305**</td>
<td>0.562**</td>
<td>0.323**</td>
</tr>
<tr>
<td>Satisfaction</td>
<td>0.666**</td>
<td>0.500**</td>
<td>0.630**</td>
<td>0.635**</td>
</tr>
<tr>
<td>MLQ</td>
<td>0.840**</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

**p ≤ 0.01 (bilateral). Source: Own work.**

**Internal consistency**

In order to evaluate the internal consistency, Cronbach’s alpha was calculated for the HSA-TFL-ES. The instrument shows satisfactory alpha levels (Cronbach’s alpha = 0.946, item-scale
correlation between 0.767 and 0.851) and we therefore assume that our results are based on reliable construct assessment.

Discussion

Although the MLQ is well known and often applied, it is too long to use in combined assessments. Therefore, a short scale for transformational leadership, based on the concept of Bass, was developed and tested. This study provides evidence for the convergent validity between the short HSA-TFL-ES and the MLQ. The results revealed a one-factor structure for the transformational leadership concept, as measured by the short HSA-TFL-ES. Furthermore, the results regarding criterion validity show that it is positively related to relevant subjective criterion variables across four nations.

Implications for theory

The results indicate that the short HSA-TFL-ES is a stable, valid and robust instrument, aspects which are important since they bring vigor to research and make the results obtained more interpretable. Although our empirical results suggest that the MLQ and the HSA-TFL-ES are highly convergent, they do show their own structure of the transformational leadership concept. The short HSA-TFL-ES appears to measure transformational leadership qualities with a unidimensional concept. Furthermore, whereas the MLQ enquires directly about the leader’s behavior, the short HSA-TFL-ES asks much more for the impact that the leader’s behavior has on the follower. This may account for the unidimensional structure of the HSA-TFL-ES, establishing the hypothesis that all transformational leadership behavior leads to follower processes. The confirmation of the unidimensionality of the HSA-TFL-ES might satisfy the need expressed by Felfe (2006) for further development of survey instruments, especially as regards newly created items, optimized scales and a simplified factor structure of the transformational leadership concept. Both approaches are related to subjective performance criteria. The short HSA-TFL-ES is notably related to relevant criterion variables across four nations, specifically commitment, identification and satisfaction. In sum, these analyses lend further support to the validity of the HSA-TFL-ES short scale.

Implications for practice

Given the needs of today’s organizations it seems necessary to bridge the gap between scientific research and practice in the organizational context (Murphy & Saal, 1990). For example, the European Foundation for Quality Management (EFQM, 2000) is dedicated to the development of organizational excellence and identifies leadership as one of its main concerns. This means that today’s organizations often need to fulfil quality requirements and, therefore, they need to combine the assessment of several human processes (Quijano et al., under revision) with the assessment of leadership and quality-related aspects of performance. Meeting these needs requires the development of instruments that are quick to apply (Carless et al., 2000), short (Burchell & Marsh, 1992) and based on scientific research (Felfe, 2006). The well-known MLQ (Bass, 1985) is rather long and this makes it difficult to use in practical circumstances. Some researchers have used a reduced set of items to measure transformational leadership (e.g. Tejeda et al., 2001), but as Rafferty and Griffin (2004) point out, this strategy has been driven by empirical results and has not been accompanied by a strong theoretical approach to explain the allocation of items to factors. They argue that it is important to adopt a theoretically driven approach when evaluating the sub-dimensions of transformational leadership. Overall, the theoretical debate and empirical results suggest that due to its relevance to today’s organizations the concept of transformational leadership should be assessable in broader practical and cultural contexts; this would be achieved by using a concept-based and valid short scale that can be combined with the assessment of other organizational and human processes.

The present results have implications for organizational assessment and intervention in these wider practical and cultural contexts. Indeed, ob-
taining an easy-to-apply, valid short instrument that measures transformational leadership goes some way to meeting the needs of enterprises for combined assessment of leadership and quality-related aspects of performance.

The unidimensionality of the construct could imply that all of the four sub-dimensions must be present for a leader to show transformational leadership. This has implications for the development of leadership in organizations, where this is a coveted quality. In this regard, an integrated leadership development program targeting all four dimensions with the same importance could be interesting.

The impact of transformational leadership as measured by the HSA-TFL-ES on subjective performance criteria such as satisfaction, identification and commitment has implications for the assessment and development of transformational leadership behavior, helping enterprises to accomplish their business goals or to manage change and improvement processes.

Limitations

The study has a number of limitations. Firstly, the analyzed samples correspond exclusively to the healthcare field and the results obtained may therefore be specific to this context.

Secondly, all measures administered were self-reported surveys. This opens up a potential problem of source bias. It can be estimated that this inflated the results in about 26% (Doty & Glick, 1998; Spector, 2006). Even if this is the case, as can be seen in table 3, considerable results still could be found.

Thirdly, we only used positive subjective performance criteria and did not include objective performance criteria or negative criteria so as to avoid common method variance (Avolio, Yammarino & Bass, 1991). Although other studies (Barling et al., 1996; Geyer & Steyerer, 1998; Rowold & Heinitz, 2007) showed an impact of TFL on objective performance criteria, it cannot be ruled out that the use of objective and negative criteria would lead to different results.

Fourthly, the subjective performance criteria used are limited and, therefore, provide limited information about the impact of transformational leadership as measured by the HSA-TFL-ES on subjective performance. Further subjective criteria such as extra-effort and effectiveness (Hetland & Sandal, 2003; Rowold & Heinitz, 2007) could produce a wider range of information about this aspect.

Fifthly, performance could be influenced by additional variables. Furthermore, context variables such as management systems or the structure of the organization may influence performance (Fuller et al., 1996; Lowe et al., 1996; Rowold & Heinitz, 2007). Thus, we cannot be sure about the extent to which leadership behavior and/or other variables contribute to performance.

Sixthly, these studies focus on followers’ ratings. In practical terms it would be interesting in the future to compare followers’ ratings and leaders’ self-ratings in a 360° feedback. Finally, TFL was measured at the same point of time, whereas longitudinal studies could better identify the possible relationship between TFL and subjective performance criteria.

Perspectives for future research

As these results were obtained in a healthcare sample, additional studies should be conducted to validate the instrument further in a different context.

Future research with the short HSA-TFL questionnaire should use hard performance criteria and unrequested outcomes such as absenteeism or experienced stress. It should also implement divergent measures in interview and observation so as to determine more clearly the criterion validity and obtain greater insight into the relationship between transformational leadership and performance. Relevant controls e.g., positive appraisal of leaders should be taken into account in future research.

The role of context variables, such as structure, should be explored in future studies to understand better their impact on performance.

The structure of the HSA-TFL construct should be analyzed using samples from other sec-
tors so as to clarify whether the results are specific for the healthcare sector.

Cross-cultural comparison analysis should be done with regard to factor structure stability. To get a deeper insight the factor structure of the HSA-TFL short should be compared cross-culturally to the MLQs factor structure.

References


Validez de la escala corta de Liderazgo Transformacional


